# Clustering

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### Overview

- Olassification (2.5 hours)
- Olustering (1.5 hours)
- Practical sessions (1 hour)

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- Understand the EM algorithm
- Being able to derive the EM updates of a mixture models

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- Understand the EM algorithm
- Being able to derive the EM updates of a mixture models
- Being able to learn by yourself!

## Outline

#### What is clustering?







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#### What is clustering?

#### 2 Mixture models

3 Admixtures











$$p(\mathbf{x}) = \sum_{k} \pi_{k} p_{\boldsymbol{\theta}_{k}}(\mathbf{x}), \qquad \sum_{k} \pi_{k} = 1, \qquad \pi \ge 0.$$

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• No closed form solution :-(







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Do we recover the original model?

$$p(\mathbf{x}) = \sum_{z} P(z)p(\mathbf{x}|z) = \sum_{k} \pi_k \text{Gaussian}(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k).$$

### Some definitions

• The differential entropy is defined as

$$H[p(\mathbf{x})] = -\int p(\mathbf{x}) \ln p(\mathbf{x}) \, d\mathbf{x}.$$

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• The Kullback-Leibler divergence measures the difference between two densities:

$$\mathrm{KL}[q\|p] = \int q(\mathbf{x}) \ln \frac{q(\mathbf{x})}{p(\mathbf{x})} \, d\mathbf{x} \ge 0.$$

The KL is asymmetric (thus not a distance) and only zero if  $q(\mathbf{x}) = p(\mathbf{x})$  for all  $\mathbf{x}$ .

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• Using *Jensen's inequality*, we get for a distribution q(Z) within a tractable family:

$$\ln p(\mathbf{x}|\boldsymbol{\theta}) = \ln \int p(\mathbf{x}, \boldsymbol{Z}|\boldsymbol{\theta}) d\boldsymbol{Z}$$
$$\geq \int q(\boldsymbol{Z}) \ln \frac{p(\mathbf{x}, \boldsymbol{Z}|\boldsymbol{\theta})}{q(\boldsymbol{Z})} d\boldsymbol{Z}$$
$$\equiv -\mathcal{F}(q, \boldsymbol{\theta}).$$

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• The quantity  $\mathcal{F}(q, \theta)$  can be interpretted as the (variational) free energy from statistical physics.

$$-\mathcal{F}(q,\theta) = \ln p(\mathbf{x}|\theta) - \mathrm{KL}[q(\mathbf{Z})||p(\mathbf{Z}|\mathbf{x},\theta)], \qquad (\mathsf{E \ step})$$
$$-\mathcal{F}(q,\theta) = \langle \ln p(\mathbf{x},\mathbf{Z}|\theta) \rangle_{q(\mathbf{Z})} + \mathrm{H}[q(\mathbf{Z})]. \qquad (\mathsf{M \ step})$$

The variational free energy  $\mathcal{F}(q, \theta)$  can be decomposed into two different ways:

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- EM can be viewed as type II maximum likelihood (ML2).

## EM in pictures



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• Maximise lower bound by alternating between:

E step: Set 
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 for fixed  $\theta$ .  
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• Gradient ascent to local maxima of  $\ln p(\mathbf{x}|\boldsymbol{\theta})$ .




### Mixture of Gaussians



$$p(\mathbf{x}|z) = \text{Gaussian}(\boldsymbol{\mu}_z, \boldsymbol{\Sigma}_z),$$
  
 $P(z) = \text{Categorical}(\boldsymbol{\pi}).$ 

• Log-complete likelihood:

$$\ln \prod_{i} p(\mathbf{x}_{i}, z_{i}) = \sum_{i} \sum_{k} \frac{\delta_{k}(z_{i})}{(\ln \pi_{k} + \ln \operatorname{Gaussian}(\boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k}))}.$$

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• Responsibilities (E step):

$$\rho_{ki} \equiv P(z = k | \mathbf{x}_i) = \frac{\pi_k \text{Gaussian}(\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_l \pi_l \text{Gaussian}(\boldsymbol{\mu}_l, \boldsymbol{\Sigma}_l)}. \quad (\star)$$













### Relation to Kmeans

• Assign data point  $x_i$  to its closest cluster:

$$r_{ki} = \begin{cases} 1 & \text{if } k = \arg \min_{l} \| \boldsymbol{x}_{i} - \boldsymbol{\mu}_{l} \|^{2}, \\ 0 & \text{otherwise.} \end{cases}$$

**Q** Recompute the cluster means after having assigned all data points.

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Let us consider  $p_{\theta_k}(\mathbf{x}) = \text{Gaussian}(\boldsymbol{\mu}_k, \boldsymbol{\epsilon} \boldsymbol{l})$ :

$$\rho_{ki} = \frac{\pi_k \exp\left(-\frac{1}{2\epsilon} \|\mathbf{x}_i - \boldsymbol{\mu}_k\|^2\right)}{\sum_l \pi_l \exp\left(-\frac{1}{2\epsilon} \|\mathbf{x}_i - \boldsymbol{\mu}_l\|^2\right)}.$$

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#### Other use cases?

Density estimation:



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### Failure mode



- PCA is a standard pre-processing tool for (linear) dimensionality reduction.
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- PPCA assumes a single Gaussian latent variable and a Gaussian likelihood.
- ML solution spans same subspace as PCA solution.
- Standard EM is  $\mathcal{O}(DNd)$  per iteration.

 $\mathbf{x}_i = \mathbf{W}\mathbf{z}_i + \mathbf{\mu} + \mathbf{\epsilon}_i$ 



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• Likelihood (noise model):

$$\mathbf{x}_i | \mathbf{z}_i \sim \text{Gaussian}(\mathbf{W} \mathbf{z}_i + \boldsymbol{\mu}, \sigma^2 \mathbf{I}_D).$$

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Likelihood (noise model):
 x<sub>i</sub>|z<sub>i</sub> ~ Gaussian(Wz<sub>i</sub> + μ, σ<sup>2</sup>I<sub>D</sub>).
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- Residual variance  $\sigma^2$  is given by  $\frac{1}{D-d} \sum_{j>d} \lambda_j$ .

### **PPCA**: interpretation



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 $p(\boldsymbol{x}) = \text{Gaussian}(\boldsymbol{\mu}, \boldsymbol{W}\boldsymbol{W}^{\top} + \sigma^2 \boldsymbol{I}_D).$ 

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- Combining local analysers to obtain nonlinear generative models.
- Possible issues are component misalignments and dimension mismatches.

# Example



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• Replace the Gaussian components by Student-t components:

$$p(\mathbf{x}) = \sum_{k} \pi_{k} p(\mathbf{x}|z=k),$$
  

$$p(\mathbf{x}|z=k) = \text{Student}(\boldsymbol{\mu}_{k}, \boldsymbol{W}_{k} \boldsymbol{W}_{k}^{\top} + \sigma^{2} \boldsymbol{I}_{D}, \boldsymbol{\nu}_{k}),$$
  

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#### Multivariate Student-t density

The Student-t density is defined as follows:<sup>1</sup>

$$\text{Student}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, \boldsymbol{\nu}) = \frac{\Gamma(\frac{\nu+D}{2})}{\Gamma(\frac{\nu}{2})(\nu\pi)^{D/2} |\boldsymbol{\Sigma}|^{1/2}} \left(1 + \frac{1}{\nu} (\boldsymbol{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})\right)^{-\frac{\nu+D}{2}}$$

Parameter  $\nu > 0$  is the shape parameter:

- The Cauchy density is recovered for  $\nu = 1$ .
- The Gaussian density is recovered when  $\nu \to \infty$ .

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The Student-*t* density can be reformulated as an infinite mixture of scaled Gaussians:

Student
$$(\mu, \Sigma, \nu) = \int_0^\infty \text{Gaussian}(\mu, \Sigma/u) \text{Gamma}(\frac{\nu}{2}, \frac{\nu}{2}) du,$$

where u is a (latent) scale parameter.

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#### Gamma density

For  $x \in \mathbb{R}^+$ , the Gamma density is defined as follows:

Gamma
$$(\alpha, \beta) = \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha-1} \exp\{-\beta x\}, \quad \alpha, \beta > 0,$$

where  $\Gamma(u) \equiv \int_0^\infty v^{u-1} e^{-v} dv$  is the gamma function.


# Example (revisited)



## USPS handwritten digits 2 and 3

- USPS data set:  $16 \times 16$  pixels images of digits (0 to 9).
- Only (respectively 731 and 658) images of digits 2 and 3 are kept.
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Standard mixture of Gaussians and diagonal mixtures collapse...





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• Each component is a product of Bernoulli distributions:

$$P_{\boldsymbol{\theta}_k}(\boldsymbol{x}) = \prod_j \operatorname{Bernoulli}(\mu_{kj}).$$



$$P(\mathbf{x}|z) = \prod_{j} \text{Bernoulli}(\mu_{zj}),$$
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• Log-complete likelihood:

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• Mean and mixture proportions (M step):

$$\boldsymbol{\mu}_k = \frac{1}{n_k} \sum_i \rho_{ik} \boldsymbol{x}_i, \qquad \pi_k = \frac{n_k}{n}, \qquad n_k = \sum_i \rho_{ik}.$$

#### Cluster means



3 components

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 $1 \ component$ 

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#### Admixtures

Mixture model:

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## Dirichlet distribution

$$\boldsymbol{\mu} \sim ext{Dirichlet} \left( \boldsymbol{\alpha} 
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• Conjugate prior to the Multinomial distribution (and Categorical):

$$p(oldsymbol{\mu}|oldsymbol{x}) \propto P(oldsymbol{x}|oldsymbol{\mu}) p(oldsymbol{\mu}) \propto \prod_j \mu_j^{x_j+lpha_j-1}$$

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ight) = rac{\Gamma(\sum_{j} \alpha_{j})}{\prod_{j} \Gamma(\alpha_{j})} \prod_{j} \mu_{j}^{\alpha_{j}-1}, \qquad \alpha_{j} \geqslant \boldsymbol{0}.$$

• Conjugate prior to the Multinomial distribution (and Categorical):

$$p(\boldsymbol{\mu}|\boldsymbol{x}) \propto P(\boldsymbol{x}|\boldsymbol{\mu})p(\boldsymbol{\mu}) \propto \prod_{j} \mu_{j}^{x_{j}+lpha_{j}-1}$$

• Defines a distribution over the simplex:







• Extremely popular (e.g., more than 14k citations in Google Scholar)



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Organise and browse large document collections



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- Capture underlying semantic structure (in an unsupervised way)
- Easily extended to discover trends, to account for the author, to model multilingual documents, to relate to the social network, etc.

#### Latent Dirichlet allocation (LDA)

#### (Blei et al., JMLR 2003)



Observations are word counts per document. LDA assumes an admix-ture model:

$$egin{aligned} \mathbf{X} \in \mathbb{N}^{V imes D}, \ \mathbf{x}_{d} \sim \prod_{i=1}^{N_{d}} \sum_{k} heta_{kd} ext{Categorical}(oldsymbol{\phi}_{k}). \end{aligned}$$

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LDA infers a low-rank approximation of the matrix of counts:

 $\mathbf{E}(\mathbf{X}) \approx \mathbf{\Phi} \mathbf{\Theta}^{\top}, \qquad \mathbf{x}_d \sim \mathrm{Multinomial}(\mathbf{\Phi} \mathbf{\theta}_d, \mathbf{N}_d)$ 

where  $\mathbf{\Phi} \in \mathbb{R}_+^{V imes K}$ ,  $\mathbf{\Theta} \in \mathbb{R}_+^{D imes K}$  and K is small.

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Simple generative model for text, based on a bag-of-words representation.

- Let V be the size of the vocabulary and K the number of topics.
- Topic k is defined as the categorical distribution  $\phi_k$  over the vocabulary.
- Document *d* is summarised as a mixture of these topics.

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- **()** The topic  $z_i$  associated to word  $w_i$  is drawn from  $\theta_d$ ; word  $w_i$  is then drawn from the categorical distribution  $\phi_{z_i}$ .

#### Graphical model and inference



 $\begin{array}{ll} \boldsymbol{\theta}_{d} \sim \mathrm{Dirichlet}\left(\alpha \mathbf{1}_{\mathcal{K}}\right), & \boldsymbol{z}_{i} | \boldsymbol{\theta}_{d} \sim \mathrm{Categorical}(\boldsymbol{\theta}_{d}), \\ \boldsymbol{\phi}_{k} \sim \mathrm{Dirichlet}\left(\beta \mathbf{1}_{V}\right), & \boldsymbol{w}_{i} | \boldsymbol{z}_{i}, \{\boldsymbol{\phi}_{k}\}_{k=1}^{K} \sim \mathrm{Categorical}(\boldsymbol{\phi}_{\boldsymbol{z}_{i}}). \end{array}$ 

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Collapsed Gibbs sampler (Griffiths and Steyvers, PNAS 2004):

$$p(z_i = k | \boldsymbol{w}, \boldsymbol{z}^{\setminus i}) \propto p(\boldsymbol{w} | \boldsymbol{z}) p(\boldsymbol{z}) \propto \frac{(\alpha + n_{\cdot kd}^{\setminus i})(\beta + n_{\vee k\cdot}^{\setminus i})}{V\beta + n_{\cdot k\cdot}^{\setminus i}},$$

where  $n_{vkd}$  is the number of times word v is assigned to topic k in document d.
"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
BEST	SPENDING	PARENTS	TEACHER
ACTOR	NEW	SAYS	BENNETT
FIRST	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
OPERA	MONEY	MEN	STATE
THEATER	PROGRAMS	PERCENT	PRESIDENT
ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

The William Randolph Heart Foundation will give 51.25 million to Lincohn Center, Metropoltan Opera Co., New York Phillamonie and Juliliad School. "Our bords field that we had a real opportunity to make a mark on the finite of the performing att with these grants an act every but as improving the second areas of support the headth, medical research, chotacian and the social services." Heart Foundation Previous Randolph A. Heart sidd Mondy in manosciencing the grant foundation of the School of the trace to building, which were the second services and the second were the performing and a set building which and the performing and a trace building which the performing area to building with a SCHOOL of the International Medical School When more of the Lincohn Center Consolidated Corporate Fund, will make its usual minimal \$100,000 domation, too.

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NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
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#### Author topic model

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#### Author topic model

#### Topics over time

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NEW	MILLION	CHILDREN	SCHOOL
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- Author topic model
- Topics over time
- N-gram topic models

"Arts"	"Budgets"	"Children"	"Education"
NEW	MILLION	CHILDREN	SCHOOL
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- Author topic model
- Topics over time
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- Author topic model
- Topics over time
- N-gram topic models
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- Multi-lingual topic models
- Topic model for images
- Population genetics
- Ο ...

# Outline

What is clustering?

2 Mixture models







# Summary

- Gaussian, Student, Bernoulli mixtures
- Alternative view of EM algorithm
- Latent Dirichlet Allocation



# Outline

What is clustering?

2 Mixture models









Derive the M step for a mixture of Gaussians.

#### References

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